Assessing the severity of emerging infections is challenging because of potential biases in case ascertainment. The first human case of infection with influenza A(H7N9) virus was identified in China in March 2013; since then, the virus has caused two epidemic waves in the country. There were 134 laboratory-confirmed cases detected in the first epidemic wave from January to September 2013. In the second epidemic wave of human infections with avian influenza A(H7N9) virus in China from October 2013 to October 2014, we estimated that the risk of death among hospitalised cases of infection with influenza A(H7N9) virus was 48% (95% credibility interval: 42–54%), slightly higher than the corresponding risk in the first wave. Age-specific risks of death among hospitalised cases were also significantly higher in the second wave. Using data on symptomatic cases identified through national sentinel influenza-like illness surveillance, we estimated that the risk of death among symptomatic cases of infection with influenza A(H7N9) virus was 0.10% (95% credibility interval: 0.029–3.6%), which is similar to previous estimates for the first epidemic wave of human infections with influenza A(H7N9) virus in 2013. An increase in the risk of death among hospitalised cases in the second wave could be real because of changes in the virus, because of seasonal changes in host susceptibility to severe infection, or because of variation in treatment practices between hospitals, while the increase could be artefactual because of changes in ascertainment of cases in different areas at different times.

Introduction
Since the first human case of infection with novel avian influenza A(H7N9) virus was identified in China in March 2013, there have been two major epidemic waves of human infections to date. The first epidemic wave, in the spring of 2013, waned during the late spring and summer [1-3], while a second major epidemic wave occurred during the winter of 2013/14 and had waned by the end of the spring of 2014 while sporadic cases have continued to be reported (as of 9 October 2014). A small number of clusters of laboratory-confirmed cases have been identified in both epidemic waves, but the virus has not appeared to have the capacity for sustained human-to-human transmission [1].

Confirmed cases of infection with influenza A(H7N9) virus have generally been identified in hospitalised patients with pneumonia [4], however, a small number of confirmed cases was identified through routine sentinel influenza-like illness (ILI) surveillance which indicates the possibility for a larger number of mild influenza A(H7N9) virus infections [5,6]. This has implications for determination of the clinical severity of influenza A(H7N9) virus infections, because the confirmed cases may not fully reflect the clinical spectrum of infections, and consequently changes in case ascertainment could lead to artefactual variation in risk of severe outcomes.

In previous work, we demonstrated that the case fatality risk among confirmed cases of infection with the 2009 pandemic influenza A(H1N1) virus was very heterogeneous and difficult to interpret [7], and we characterised the severity of influenza A(H7N9) virus infections via the risk of fatalities among hospitalised cases (the 'hospitalisation fatality risk', HFR) and the risk of fatalities among symptomatic cases (the 'symptomatic case fatality risk', CFR) [3]. In the first epidemic wave of influenza A(H7N9) virus infections in spring 2013, we estimated the HFR at 36%, and the CFR at
The first wave of infections in 2013 is divided into two parts, before and after the announcement of human cases on 31 March 2013 because of the potential for under-ascertainment of less severe cases in the earlier period.
number of mild cases as the denominator. We used a Bayesian framework to estimate the symptomatic CFR, and presented the estimates with 95% credibility intervals (CrI) which have a similar interpretation to confidence intervals [9].

We examined epidemiologic time-to-event distributions using kernel density methods as previously described [2]. All statistical analyses were performed using R version 3.0.1 (R Foundation for Statistical Computing, Vienna, Austria) and Matlab (Mathworks Inc., Natick, Massachusetts, United States).

Results

In the first wave of influenza A(H7N9) cases in 2013, 134 confirmed cases were identified (Figure 1), of whom 124 required hospitalisation for medical reasons. Among the hospitalised cases, the risk of serious outcomes was higher among older hospitalised cases. Furthermore, we identified higher risks of fatalities among cases hospitalised before 31 March 2013, the date when the first confirmed human cases of influenza A(H7N9) virus infection were officially announced in China (Figure 2).

We therefore divided the first wave into two parts: wave 1A for 18 cases hospitalised before 1 April 2013, and wave 1B for 106 cases hospitalised from 1 April to 30 September 2013 (Figure 1).

In the first epidemic wave, the median age was 60 years in wave 1A and 61 years in wave 1B. Among the cases under 60 years who required hospitalisation for medical reasons, the HFR in wave 1A was 51% (95% CI: 21%–79%), significantly higher (p = 0.039) than the HFR of 17% (95% CI: 7.6%–30%) in wave 1B. For cases above 60 years who required hospitalisation for medical reasons, the HFR was also significantly higher (p = 0.025) in wave 1A (77%; 95% CI: 48%–94%) vs wave 1B (42%, 95% CI: 31%–54%). We did not identify significant differences between wave 1A and 1B in the risk of death or ventilation, or in the risk of death/ventilation/ICU admission (Figure 2).

In the second epidemic wave of influenza A(H7N9), 273 of the 306 confirmed cases required hospitalisation for medical reasons with onset dates between 1 October 2013 and 9 October 2014. The median age was 57 years (range 2–88 years). Sixty-nine percent of cases were male. Among the hospitalised cases, allowing for censoring of outcomes in five (2%) patients remain in hospital on 9 October 2014, we estimated HFRs of 36% (95% CI: 28%–45%) in cases under 60 years, and 59% (95% CI: 51%–67%) in cases aged 60 years or above. These risks were significantly higher than in wave 1B (p = 0.019 and p = 0.025 respectively). There were no statistically significant differences between the age-specific risks of death or ventilation, or death/ventilation/ICU admission in wave 2 compared to either wave 1A or wave 1B, while estimates of the risks of serious outcomes were generally lower across age groups in wave 1B compared with wave 2 (Figure 2).
therefore examined the risk of death among the subset of hospitalised cases in this province. Zhejiang province reported 40 cases in wave 1B and 88 cases in wave 2, and the risk of death among hospitalised cases under 60 years-old was significantly higher in wave 2 compared with wave 1B (risk ratio 7.1; 95% CI: 1.3–292; p = 0.017) and not significantly different in hospitalised cases above 60 years-old (risk ratio 1.5; 95% CI: 0.93–2.8; p = 0.099).

We examined the delays from onset to admission and identified similar patterns over calendar time, while the delay from onset to laboratory confirmation has shortened over time and in wave 2 the mean was eight days (Figure 3). Distributions of time from admission to death and from admission to discharge were similar over time (Figure 3).

We previously used information on three confirmed influenza A(H7N9) cases identified through ILI surveillance in Shanghai and Nanjing to estimate the number of symptomatic cases in the spring 2013 epidemic wave [3]. Here we also use information on four confirmed cases identified through ILI surveillance in Shaoxing in the winter 2013/14 epidemic wave, in the period from 1 January to 21 January 2014, before the closure of live poultry markets on 22 January. During the same period in Shaoxing, nine hospitalised cases had onset of illness, of whom five died. Based on these observations, we estimated that there were 3,020
cases in the first epidemic wave in 2013 in Shanghai and Nanjing, respectively, and 5,750 (95% CI: 1,960–12,730) cases in Shaoxing in the second epidemic wave in 2013/14. These estimates correspond to symptomatic CFRs of 490 and 69 in Shanghai and Nanjing respectively in the first wave, and 100 per 100,000 symptomatic cases in Shaoxing in the second wave, with wide and overlapping credibility intervals (Table).

Discussion

The resurgence of human infections with avian influenza A(H7N9) virus in a second epidemic wave in 2013/14 demonstrates the continued public health risk of this novel strain [10]. Control of the virus in animals is complicated, because the infections in poultry are asymptomatic [11]. Human-to-human transmissibility of the virus remains limited, as evidenced by the very small number of potential secondary infections identified through detailed contact tracing of confirmed cases [1,2,12-14].

We identified differences in the severity of illness of hospitalised cases in the earlier part of the first epidemic wave in 2013, with greater risk of mechanical ventilation, ICU admission and death among cases hospitalised before 31 March 2013 when the first confirmed human cases of influenza A(H7N9) were officially announced (Figure 2) [15]. One explanation for this is more timely antiviral treatment and more appropriate supportive care for cases hospitalised after 31 March 2013. Another possible explanation is detection bias in the early phase of the spring 2013 epidemic wave, where more severe cases were prioritised for repeated laboratory testing, and cases with prolonged virus shedding or higher virus shedding had a greater chance of confirmation.

In the second epidemic wave in 2013/14, we identified a significantly greater HFR compared with the latter part of the first epidemic wave in 2013 (Figure 2) and in persons under 60 years of age in Zhejiang province where cases occurred in both epidemic waves, but no difference in the symptomatic CFR (Table). It is possible that this significant difference in HFRs is due to ascertainment bias in cases in different locations at different times, even within the same province. Alternatively, the HFR could have increased, because hospitalised cases in the second epidemic wave in 2013/14 were less likely to be transferred to larger referral hospitals (Dr Enfu Chen, Chief Epidemiologist in Zhejiang Provincial CDC, personal communication, June 2014), because of changes in the virus, or because of seasonal changes in the prevalence of other pathogens that could cause secondary or co-infections and modify the severity of influenza A(H7N9) virus infections [16]. Whereas ascertainment of infections in hospitalised cases may have changed over time due to changes in awareness and testing capacity, the ascertainment of influenza A(H7N9) cases through the established sentinel ILI network should have remained more stable over time.

Large population-based serological studies in affected areas would permit assessment of severity with a denominator of infections, rather than cases of symptomatic disease or hospitalisation, and infection-based severity measures could be less susceptible to biases due to differential healthcare seeking behaviours or diagnostic capacity [3,7]. To date, few serological studies have been reported and such analyses are not yet possible [17-19].

Our estimates of the risks of serious outcomes in hospitalised cases are limited by the potential for under-ascertainment of cases, due to lack of access to laboratory testing in some areas, and the potential for imperfect sensitivity of laboratory testing for the A(H7N9) virus [20,21]. While we accounted for unknown final status of cases that remain hospitalised in our analysis, the eventual estimates may change slightly once all outcomes are known. It is challenging to estimate the symptomatic CFR based on a small number of confirmed cases with milder disease identified through sentinel ILI surveillance, and our estimates are dependent on the assumptions that coverage of the sentinel system was similar in 2013/14 compared with 2009, and that healthcare seeking behaviours for ILI were similar whether illness was caused by influenza

<table>
<thead>
<tr>
<th>Period analysed</th>
<th>Geographic location</th>
<th>Number of confirmed deaths caused by influenza A(H7N9) virus infection</th>
<th>Estimated number of symptomatic A(H7N9) virus infections</th>
<th>Estimated risk of fatalities per 100,000 symptomatic cases</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Jan 2013–28 May 2013</td>
<td>Shanghai</td>
<td>14</td>
<td>3,020 (95% CI: 900–7,800)</td>
<td>490 (95% CI: 170–1,800)</td>
</tr>
<tr>
<td>1 Jan 2013–28 May 2013</td>
<td>Nanjing (Jiangsu province)</td>
<td>3</td>
<td>5,310 (95% CI: 880–17,300)</td>
<td>69 (95% CI: 12–710)</td>
</tr>
<tr>
<td>1 Jan 2014–21 Jan 2014</td>
<td>Shaoxing (Zhejiang province)</td>
<td>5</td>
<td>5,750 (95% CI: 1,960–12,730)</td>
<td>100 (95% CI: 29–360)</td>
</tr>
</tbody>
</table>

CI: confidence interval.
A(H7N9) virus or the 2009 pandemic influenza A(H1N1) virus [3]. In addition, the estimation of scFR were based on data from geographic locations in which influenza A(H7N9) virus infections were identified through sentinel ILI surveillance, and a more comprehensive analysis could also incorporate data on ILI surveillance in other areas.

In conclusion, it remains important to assess the severity of human infections with influenza A(H7N9) virus, as part of ongoing risk assessment of this virus. While the overall picture is that the severity of human infections has not substantially changed (Table), we found some evidence that the HFR was higher in the second epidemic wave in 2013/14 (Figure 2). Our results again highlight that many influenza A(H7N9) virus infections can cause mild disease [3,5,6] and that the risk of death among laboratory-confirmed cases is a misleading measure of severity. If another epidemic of human infections with influenza A(H7N9) virus occurs in the winter of 2014/15, proactive control measures on the poultry-human interface may be preferable to reactive measures [10,22-24]. Comprehensive surveillance of avian influenza virus infections in animals and humans is essential in order to monitor risk and guide the use of control measures.

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Conflicts of interest
GML has received consulting honoraria from Janssen Pharmaceuticals. BJC reports receipt of research funding from MedImmune Inc. and Sanofi Pasteur, and consults for Crucell NV.

Authors’ contributions
Hongjie Yu and Benjamin J Cowling designed the study. Luzhao Feng, Joseph T. Wu, Xiaoxing Liu, Peng Yang, Tim K. Tsang, Huijiang Peng, Wu Fu, Yuan Jiang, Vicky J. Fang, Ying Qin, Eric H. Y. Lau, Ming Li, Jianhong Zheng, Zhbin Peng, Yun Xie, Quanyu Wang, Zhongjie Li, Guibing Li, Fei Qiao and George F. Gao collected data. Luzhao Feng, Joseph T. Wu, Tim K. Tsang, Peng Wu, Vicky J. Fang and Eric H. Y. Lau analysed data. Benjamin J Cowling wrote the first draft and all authors contributed to review and revision and have seen and approved the final version.

References


